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Language learning strategies: The adaptation of the Czech-language version of the Strategy Inventory for Language Learning for secondary school students

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Abstract: The Strategy Inventory for Language Learning (SILL) is a common tool for assessing language learning strategies. However, due to changes in learning environments, such as increased digital exposure and out-of-class language practice, the original structure may not fully capture current learner strategies. Thus, validation is still needed. This study details the adaptation process of the questionnaire for the Czech context. It involved 1306 secondary school students (706 female, 555 male, and 45 identifying differently). Both exploratory and confirmatory factor analyses were conducted, testing two-, three-, and four-factor models. The two-factor model (CFI = .918, TLI = .901, RMSEA = .066, CMIN/DF = 3.806) provided the best fit. The final version includes both direct and indirect strategies. The tool shows excellent internal consistency (Cronbach's $\alpha = .861$; McDonald's $\omega = .853$) and is useful for diagnostics and self-monitoring.

Keywords: Strategy Inventory for Language Learning, language learning strategies, adaptation, exploratory factor analysis, confirmatory factor analysis, reliability

Strategie učení se jazykům: Adaptace české verze dotazníku Strategy Inventory for Language Learning pro žáky středních škol

Abstrakt: Strategy Inventory for Language Learning (SILL) je široce využívaný nástroj pro měření strategií učení se jazykům. V důsledku proměn podmínek jazykového učení, včetně rostoucí expozice digitálních technologií a mimoškolního kontaktu s jazykem, však původní struktura nástroje nemusí plně odrážet strategie využívané současnými žáky. Cílem této studie je popsat proces adaptace dotazníku pro českou populaci. Výzkumu se zúčastnilo 1306 středoškolských žáků (706 dívek, 555 chlapců a 45 žáků uvedlo jiné pohlaví). Pro účely adaptace byly použity explorační a konfirmační faktorová analýza. Testovány byly dvou-, tří- a čtyřfaktorové modely. Na základě porovnání ukazatelů shody (CFI = 0,918, TLI = 0,901, RMSEA = 0,066, CMIN/DF = 3,806) se jako nejvhodnější ukázal dvoufaktorový model SILL-CZ. Finální verze nástroje zahrnuje přímé a nepřímé strategie. Nástroj vykázal velmi dobrou vnitřní konzistenci (Cronbachovo $\alpha = 0,861$; McDonaldovo $\omega = 0,853$) a může sloužit jako praktický diagnostický nástroj i prostředek pro sebemonitorování učení.

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Klíčová slova: strategický inventář pro výuku jazyků, strategie učení se jazykům, adaptace, explorační faktorová analýza, konfirmační faktorová analýza, spolehlivost

1 Introduction

The European Union Council (2018) highlights multilingualism as one of eight essential competencies for promoting a healthy, sustainable lifestyle, employability, civic engagement, and social inclusion. Proficiency in foreign languages is becoming more crucial for both work and social interactions (Council of the European Union, 2019). In the EU, foreign language instruction begins at early education levels (EACEA, 2023).

Language Learning Strategies (LLS) are essential in learning a foreign language, as they promote more effective, autonomous, and self-regulated learning. Studies link LLS to better outcomes, increased motivation, and greater learner independence. Students who use a broader range of strategies often achieve higher success (Sukyng, 2021; Wong, 2011; Yoong & Hashim, 2023). Additionally, the rise of digital technologies and exposure to English outside the classroom has further expanded learning opportunities (Jahrani & Listia, 2023; Sundqvist, 2011; Viberg & Kukulska-Hulme, 2021). Effective strategies are beneficial not only for high-achieving students; less successful learners can also improve by adopting and developing them (Alfarisy, 2022). Since strategy use varies among learners, it is important to identify which strategies are favoured or neglected and how their use evolves over time and contexts (Chong et al., 2025; Papadopoulou et al., 2018; Yoong & Hashim, 2023). The Czech context might display specific patterns of strategy use due to cultural, educational, and curricular differences. Secondary education is a vital stage for fostering more effective learning habits, autonomy, and responsibility for learning (Weinstein & Acee, 2013). Therefore, a tool is needed to measure and assess the LLS used by Czech students. The most common instrument in this area is the Strategy Inventory for Language Learning (SILL) by R. Oxford (Domínguez & Juanías, 2024), and this study focuses on adapting it for the Czech context.

The SILL has been previously studied within the Czech context (Vlčková, 2007; Vlčková & Přikrylová, 2011). Past research has clarified the factor structure of language learning strategies and led to an adapted version suited to the Czech educational system. As language learning conditions continue to evolve and the demand for relevant instruments grows, ongoing validation of the SILL remains important.

Considering the points above, it is important to examine how language learning conditions are evolving today. Although Language Learning Strategies (LLS) are well documented in existing literature, the relevance of the SILL instrument remains given recent changes. These include the growing influence of digital technologies, increased extramural exposure to English, and shifts in upper secondary education, such as updates to the school-leaving exam format. From a social pedagogical standpoint, understanding how students handle classroom demands and approach their learning is vital, as these factors are closely linked to academic achievement and potential school failure. Consequently, learning strategies are seen not only as a didactic concept but also as a reflection of how students organise, regulate, and manage their learning processes.

The validated instrument is useful not only for teachers but also for school social pedagogues who assist students in at-risk situations. It helps identify learning gaps and promotes targeted development of effective strategies, thereby aiding the prevention of school failure through timely, individualised support.

2 Theoretical background

Research on Language Learning Strategies (LLS) has spanned over fifty years, evolving in both theoretical foundations and the key concepts involved (Pawlak, 2021). Since the 1960s, the focus has shifted from the teacher and teaching methods to the learner's learning process, reflecting a move from behaviourist to cognitivist paradigms (Vlčková, 2010), which laid the groundwork for LLS research. Numerous definitions and classifications of LLS have been proposed. Oxford (1989) described LLS as "behaviours or actions which learners use to make language learning more successful, self-directed and enjoyable". In 1990, she provided another definition, describing them as "specific actions taken by the learner to make learning easier, faster, more enjoyable, more self-directed and more transferable to new situations" (Oxford, 1990). O'Malley and Chamot (1990) define LLS as "special thoughts or behaviours that individuals use to help them comprehend, learn, or retain new information" (p. 1). Despite differences, both highlight the learner's active participation in enhancing language learning. This perspective informs this study, especially in adapting the research instrument.

Numerous assessment tools for LLS have been developed over the years (Charoento, 2017; Cohen, 1990; Radwan 2011; Sánchez, 2019; Teng & Zhang, 2016; Tseng et al., 2006; Wang et al., 2013; Zhang & Seepjo, 2013). This classification also mirrors the development of research and understanding of LLS, which is not uniformly defined. It is important to recognise that experts from various disciplines work on this topic. Consequently, certain criteria might be integrated into the concept of strategies to inform research tools. Therefore, strategies can be categorised based on their environment or particular perspectives, such as self-regulation (Habók et al., 2022), speech skills and language resources (Cohen & Weaver, 2006; Oxford, 1990), stages of the learning process (Bimmel & Rampllion, 2000), and others.

Despite the variety, the SILL developed by R. L. Oxford has remained prominent in the field for decades (Domínguez & Juanías, 2024; Rose et al., 2018; Saks & Leijen, 2016). This tool was created in two versions: one with 80 items for native English speakers, and a 50-item version for learners for whom English is a foreign language. The 50-item version is mainly divided into direct and indirect strategies. Direct strategies include memory (9 items), cognitive (14 items), and compensatory strategies (6 items). Indirect strategies encompass metacognitive (9 items), affective (6 items), and social strategies (6 items). This classification has faced criticism. For example, Dörnyei (2005) argues that compensatory strategies relate more to language use than learning. He also suggests that memory strategies should be a subset of cognitive strategies. Dörnyei views strategy terminology sceptically and proposes replacing it with 'self-regulation,' which he believes covers a broader range of processes and skills. Other researchers have expanded on this view, emphasising the importance of self-regulation in language learning (Dörnyei, 2005; Habók et al., 2022). Conversely, some researchers continue to treat Language Learning Strategies as a distinct concept (Oxford, 1990; Vlčková, 2007). Oxford (2017) later incorporated self-regulation into her framework while reaffirming the ongoing relevance of LLS as a concept.

LLS has faced various criticisms and challenges, yet its ability to understand second language learning remains clear (Pawlak, 2021). A systematic review by Domínguez and Juanías (2024) highlights the ongoing prominence of the LLS concept, largely attributed to R. L. Oxford, the most-cited author in the field. Oxford's (1990) classification provided the theoretical foundation for this research for several reasons: it is detailed and comprehensive, including all known basic strategies; it can be practically applied in teaching for diagnosis and intervention; other researchers also utilize its extensive item bank for developing their own tools (Habók et al., 2022; Vlčková, 2010); and it employs fewer abstract or complex terms compared to other classifications based on linguistic and psychological theories.

The SILL is a widely used tool globally, as evidenced by its application across continents. Oxford (1996) noted that the use of LLS varies according to different educational and cultural backgrounds. Most research focuses on examining English learning strategies among various nations, including China (Deneme, 2008; Hu, 2010), South Korea (Park, 2011), Taiwan (Wu, 2008), Singapore (Gu et al., 2005; Wharton, 2000), Japan (Robson & Midorikawa, 2001), Turkey (Yesilbursa & Ipek, 2013), Greece

(Petrogiannis & Gavriilidou, 2015), and Peru (Saez-Zevallos et al., 2025). Other studies have explored intercultural differences in language learning strategies (Banasiak-Ryba, 2010; Deneme, 2010; Habók et al., 2021; Hong-Nam & Leavell, 2007; Jiang & Wu, 2016; Oxford, 1996). Since its development, the SILL has been adapted in numerous cultural contexts, but results have varied (Park, 2011). Some researchers validated the original six-factor structure (Demirel, 2009; Wu & Chang, 2018), while others did not (Park, 2011; Yesilbursa & Ipek, 2013). Park (2011) highlighted that educational and cultural contexts influence how SILL items group into factors, arguing that the original six-factor model is not universally applicable. Although he tested the original structure, his analysis ultimately suggested a simpler model with fewer factors. Conversely, Petrogiannis and Gavriilidou (2015) tested both six- and two-factor models (reducing the instrument to 29 items), both showing good internal consistency. Robson and Midorikawa (2001), through Exploratory Factor Analysis (EFA), identified 15 new strategic areas from the SILL, but these factors had weak internal consistency, leading them to regard the SILL as a general measure of language learning strategies rather than a tool for six specific types. Past research has also examined connections between individual language skills and learning strategies, especially the links between reading, writing, and speaking (Ansarey, 2016; Esmaeili, 2002; Jia, 2023; Namanziandost et al., 2018), including how reading and writing support oral language development (Ergün Öcel, 2023). Additionally, media exposure has been considered influential in language development (Ennemoser & Schneider, 2007). These findings suggest learning strategies are closely tied to broader language use and skill acquisition. Recently, the context of language learning has shifted significantly, particularly due to increased exposure to digital media and multimodal communication, which may influence strategy use and raise questions about the ongoing relevance of instruments such as the SILL.

This study aims to adapt the SILL version of the "questionnaire for speakers of other languages learning English" for Czech secondary school students. Because of this, revalidating the instrument in the Czech context under current educational conditions is necessary. This is especially important in upper secondary education, where students must pass the State Matura Examination. In its general part, students are required to pass Czech language and literature and choose between a foreign language and mathematics. Since the 2021/2022 school year, this choice has been mandatory, with about 80% of students opting for English as their second Matura subject in 2024. Meanwhile, the preference for mathematics has been steadily decreasing over the past decade (43.9% in 2012; CERMAT, 2025).

3 Research design

3.1 Participants

The survey included 1,306 students from 17 randomly selected vocational secondary schools and grammar schools across the Czech Republic. Of these, 706 were female, 555 male, and 45 identified with another gender. The average age was approximately 17 years (range: 15–20). The sample represents the entire Czech Republic. An online questionnaire, distributed via Google Forms, was sent to school principals after they agreed to participate. Participants saw an informed consent form arranged by the schools. The principals informed students about voluntary participation, the study's purpose, understanding of the information, and how to complete the questionnaire. No questionnaires were excluded. In designing the survey, each item was mandatory, preventing submission unless all items were answered. Data collection occurred during lessons, in accordance with each school's preference. The study aimed to minimise any risk to participants. Data were collected through standardised questionnaires completed online without researchers present. The study's potential benefits are significant: adapting the instrument may help measure Language Learning Strategies in the Czech population, aiding research and teaching. Data collection took place from autumn 2022 to spring 2023. The research was conducted ethically and received approval from the University's Ethics Committee.

3.2 Research instrument

The study used the SILL Version for Speakers of Other Languages Learning English (Oxford, 1989) as its primary tool. This inventory contains 50 items, mainly categorised into direct and indirect strategies. Direct strategies comprise memory (9 items), cognitive (14 items), and compensatory strategies (6 items). Indirect strategies include metacognitive (9 items), affective (6 items), and social strategies (6 items). Participants responded on a five-point Likert scale, from 1 (“always or almost always false”) to 5 (“always or almost always true”).

3.3 Research procedure

The adaptation of the SILL-CZ questionnaire consisted of three phases. Initially, the original SILL version for speakers of other languages learning English (Oxford, 1989), which contains 50 items and was used in this study, was translated into Czech. Two independent translators performed the translation, which was then evaluated by two experts through semantic editing to ensure the content standards were preserved. After this revision, the questionnaire was prepared for subsequent stages of adaptation. The second phase involved a pilot study with secondary school students to assess the clarity of the questions and overall usability, including the statements and rating scale. Data were collected from 102 students across one of the fourteen Czech regions; these students were excluded from later phases. No significant changes were made during this stage, and all items were deemed clearly understandable. The third phase focused on further adapting the instrument using the analytical procedures described below.

3.4 Data analysis

The adaptation of the SILL-CZ research tool was conducted in two stages. Initially, Exploratory Factor Analysis (EFA) was used, followed by Confirmatory Factor Analysis (CFA). EFA is usually applied when developing a new instrument, aiming to identify unknown latent factors that best fit the data's correlation structure (Field, 2024). Conversely, CFA tests a predefined model of expected relationships to confirm or refute it based on the data (Brown, 2006; Field, 2024; Finch, 2020; Soukup, 2022). Oxford and Amerstorfer (2018) highlight that although SILL is significant in Language Learning Strategies (LLS), it should be adapted to specific research needs. Both EFA and CFA were utilized here because adapting an instrument for use across different countries and cultures necessitates employing both analytical methods.

The data were analysed using IBM SPSS Statistics for descriptive statistics, EFA, and reliability tests, and IBM SPSS Amos for CFA. The dataset from 1,306 participants was divided into two subsets with the Random Sample of Cases procedure to avoid using the same data for both EFA and CFA, as this would violate standard psychometric practices (Brown, 2006; Soukup, 2021). EFA was performed on 635 cases, while CFA was conducted on 671. The adaptation questionnaire comprised 50 items rated on a five-point scale (1 - always/almost always false; 5 - always/almost always true). Because the scale is longer and ordinal, Pearson correlations were used in EFA instead of polychoric correlations (Soukup, 2021). Both EFA and CFA tested three models: two-, three-, and four-factor solutions. The best-fitting model, according to validity and reliability criteria, proceeded to further analysis. Internal consistency was assessed with both Cronbach's α and McDonald's ω .

4 Results

4.1 Exploratory factor analysis

Before conducting the EFA, it was necessary to verify whether the data were suitable for factor analysis. For this purpose, the Kaiser–Meyer–Olkin (KMO) measure, Bartlett's test, and the Measure of Sampling Adequacy (MSA) were used. The Kaiser–Meyer–Olkin (KMO) statistic is a commonly used measure in EFA to assess whether the data are suitable for factor analysis. It quantifies the sampling

adequacy of each observed variable in the model, as well as the overall model. The KMO ranges from 0 to 1. The closer the value is to 1, the more suitable the data are for EFA. The Measure of Sampling Adequacy (MSA) evaluates the appropriateness of each variable for inclusion in factor analysis. The overall KMO statistic represents the average of all individual MSA values (Kaiser, 1974). Bartlett's test examines whether the correlation matrix among variables differs significantly from the identity matrix. If the correlation matrix were an identity matrix, the variables would be uncorrelated, and factor analysis would not be appropriate (Dyer & Keating, 1980). This test evaluates the null hypothesis that the correlation matrix is an identity matrix. If $p < .05$, the null hypothesis is rejected, indicating that the data are suitable for factor analysis.

All three-factor solutions—two-, three-, and four-factor—met the criteria for factorability, with values within acceptable ranges (Mareš et al., 2015). Bartlett's test p -value was $<.001$, the KMO was .939, and MSA values remained above .70 for all items. Communalities (h^2) were checked prior to factor extraction, representing the proportion of each item's variance explained by the common factors (Tavakol & Wetzel, 2020). The Fixed Number of Factors method was employed because the scree plot indicated several factors with only two items each, which is inadequate for factor development (MacCallum et al., 1999; Raubenheimer, 2004). Additionally, an Absolute Value Below .10 was used to identify factor loadings across multiple factors, facilitating the CFA process. Principal Axis Analysis (PA), suitable for non-normal data, was used in EFA to uncover latent factors that explain shared variance among variables (Finch, 2020). Oblimin rotation, which allows for correlated factors (Field, 2024), was applied. The analysis followed the .40-.30-.20 rule, as described by Howard (2016, p. 59), stating: "This rule recommends that only variables that (a) load above .40 on their primary factor, (b) cross-load below .30 on other factors, and (c) show a difference of at least .20 between their main and other loadings are interpreted." Applying this rule reduced the number of items in the EFA, resulting in the different item counts for the two-, three-, and four-factor solutions (see Supplementary Tables 1 and 2). The full EFA results for the two-factor solution are shown in Table 1.

All three models underwent exploratory factor analysis (EFA) and internal consistency testing using appropriate coefficients. They all satisfied the conditions for EFA, including KMO, Bartlett's test, and MSA. However, examining the factor loadings revealed differences across models. The first factor of the two-factor model showed very satisfactory loadings (maximum = .802, minimum = .443). The second factor also had satisfactory loadings (maximum = .625, minimum = .441), despite containing only six items. The three-factor model generally did not exhibit low factor loadings, except for the third factor, where most items failed to reach the .40 threshold. The four-factor model performed as well as the two-factor model but demonstrated lower internal consistency for its last two factors. Despite having 21 items, the two-factor model displayed the highest overall internal consistency ($\alpha = .899$, $\omega = .899$) and in its individual factors (model F1: $\alpha = .919$, $\omega = .709$; model F2: $\alpha = .921$, $\omega = .710$).

4.2 Confirmatory factor analysis

EFA was conducted on the first half of the dataset, while the second half was used for CFA. Basic descriptive statistics, including minimum, maximum, mean, standard deviation, skewness, and kurtosis, were computed. The multivariate normality test indicated a non-normal data distribution. Several methods exist for estimating the common factor model; among them, Maximum Likelihood (ML) is common (Brown, 2006). In this study, the ML estimation with the Bollen-Stine Bootstrap correction was selected. Although ML is optimal for larger samples and assumes normality, the data's non-normality necessitated the use of the Bollen-Stine Bootstrap to address outliers and distributional issues (Ward & Ahlquist, 2018). CFA was first performed on each factor independently, providing insights into the instrument's internal structure and covariances, including errors. Items with standardised factor loadings below .40 were removed (Stevens, 2009). The final model retained items with loadings between .46 and .76. Examination of Standardised Residual Covariances confirmed all models were adequately reduced in item number.

Table 1
Factor loadings and communalities for the two-factor solution in the EFA of the SILL.

Item	Factor		h^2
	1	2	
I1 I try to find as many ways as I can to use my English.	.802		.644
I2 I try to talk like a native English speaker.	.772		.597
I3 I ask questions in English.	.748		.559
I4 I start conversations in English.	.746		.563
I5 I write notes, messages, letters, or reports in English.	.711		.534
I6 I read English for pleasure.	.709		.564
I7 I look for opportunities to read as much as possible in English.	.687		.473
I8 I watch English-language TV shows or go to English-language movies.	.679		.505
I9 If I can't think of an English word, I use a word or phrase that means the same thing.	.674		.463
I10 I read English without looking up every new word.	.650		.496
I11 I encourage myself to speak English even when I am afraid of making a mistake.	.618		.388
I12 I try to learn about the culture of English speakers.	.582		.342
I13 I try not to translate word for word.	.497		.248
I14 To understand unfamiliar English words, I make guesses.	.464		.224
I15 I try to guess what the other person will say next in English.	.443		.199
I16 I plan my schedule so that I have enough time to study English.		.625	.398
I17 I try to find out how to be a better learner of English.		.610	.405
I18 I review English lessons regularly.		.594	.354
I19 I look for words in my own language that are similar to new English words.		.510	.264
I20 I give myself a reward or treat when I do well in English.		.442	.204
I21 I say or write new English words several times.		.441	.195
			Total
Cronbach's α	.919	.709	.899
McDonald's ω	.921	.710	.899

Note. Extraction method: Principal Axis Factoring. Rotation Method: Oblimin with Kaiser Normalisation. I = item; h^2 = communalities.

In CFA, the model was modified to include covariances among item error terms, as previously mentioned. The low covariance of some items with the factor might be because these items share something in common that others with high loadings do not (Soukup, 2022). There were six pairs of linked items: five related to the first factor and one to the second factor. The weakest correlation was between items 14 and 17 ($r = .176$). Table 2 presents all covariances between the error terms of the items.

CFA was used to assess the values in the Standardised Residual Covariances section. 4 items did not meet the criteria and were therefore excluded from the next stage of the analysis. The recommended threshold is 2.58 (Jöreskog & Sörbom, 1993). Factor 1 (Direct Strategies) comprised 11 items, and Factor 2 (Indirect Strategies) comprised 6 items.

Various analyses were conducted during the CFA to assess whether the factors identified after the EFA matched the expected structure. Table 3 compares the two-factor model with alternative models and presents the results. In SEM, multiple fit indices are available, but researchers have not reached a consensus on which to prioritise (Bhale & Bedi, 2023). This study employed the Comparative Fit Index (CFI), Tucker-Lewis Index (TLI), and Root Mean Square Error of Approximation (RMSEA), following Xia and Yang (2019), to evaluate model fit. The CFI measures how well the model fits the observed data.

The TLI compares the proposed model to a null model that assumes no relationships among variables, while adjusting for model complexity. The RMSEA assesses how closely the model approximates the data per degree of freedom, while also accounting for complexity. Bentler (1990) recommends a CFI above .90, while Bentler and Bonnet (1980) suggest a TLI above .90. Browne and Cudeck (1993) set an RMSEA threshold at less than .08. Wheaton et al. (1977) mention that a CMIN/DF ratio of ≤ 5 indicates a reasonable fit, where CMIN is the chi-square discrepancy, and df is degrees of freedom.

Table 2

Covariances between item error terms in the two-factor CFA model of the SILL.

Factor	Item
F1	I5: I write notes, messages, letters, or reports in English. I6: I read English for pleasure.
F1	I5: I write notes, messages, letters, or reports in English. I9: I encourage myself to speak English even when I am afraid of making a mistake.
F1	I6: I read English for pleasure. I7: I look for opportunities to read as much as possible in English.
F1	I6: I read English for pleasure. I9: I encourage myself to speak English even when I am afraid of making a mistake.
F1	I6: I read English for pleasure. I8: I watch English-language TV shows or go to English-language movies.
F2	I14: I review English lessons regularly. I17: I say or write new English words several times.

Note. I = item; F1 = Factor 1; F2 = Factor 2.

The results indicated that the two-factor model of the 17-item questionnaire is the most suitable for the data, as it fits significantly better than the three- and four-factor models. This is supported by the indices, which show the best values for the two-factor model (CFI = .918, TLI = .901, RMSEA = .066, CMIN/DF = 3.806).

Table 3

Comparative fit indices for two-, three-, and four-factor CFA models

Model	No. of items	CMIN/DF	CFI	TLI	RMSEA
Two-factor	17	3.806	.918	.901	.066
Three-factor	16	3.972	.910	.888	.068
Four-factor	24	4.369	.881	.861	.067

Note. No. = number of items; CMIN/DF = Minimum variance of CFA/degrees of freedom; CFI = The comparative fit index; TLI = Tucker-Lewis index; RMSEA = Root mean square error of approximation.

4.3 Two-factor model of SILL-CZ

Supplementary Table 3 displays descriptive statistics for all items in the instrument. The mean scores for the final 17 items ranged from 1.95 to 3.81 on a 1-5 scale, with standard deviations (SD) between 1.110 and 1.515. None of the items showed skewness or kurtosis exceeding ± 2 (Trochim & Donnelly, 2006). The next step involved calculating correlations to assess the validity of the two-factor model. A moderate correlation of $r = .344$ was found between Factor F1 (Direct Strategies) and Factor F2 (Indirect Strategies). As Cohen (1988) classifies, low correlation is around $r = .10$, medium around $r = .30$, and high around $r = .50$. Supplementary Figure 1 illustrates the two-factor solution model.

4.4 Final internal consistency of the SILL-CZ

The final stage of data analysis focused on determining the instrument's internal consistency. Comparisons were again made between the predicted and alternative models, with results shown in

Table 5. It is evident that only the two-factor model met the internal consistency criterion for coefficient alpha, exceeding the .70 threshold, indicating that the items are internally reliable (Mareš et al., 2015). Specifically, for factor 1, $\alpha = .885$ and $\omega = .887$ were observed, whereas for factor 2, $\alpha = .748$ and $\omega = .748$. In the three-factor model, only the first factor exceeded the threshold, and in the four-factor model, the first two factors met the criteria.

Table 5

Final internal consistency for the two-, three-, and four-factor model of the SILL-CZ.

Model	Cronbach's α					McDonald's ω				
	F1	F2	F3	F4	Total	F1	F2	F3	F4	Total
Two-factor	.885	.748			.861	.887	.748			.853
Three-factor	.898	.601	.546		.804	.848	.591	.543		.780
Four-factor	.929	.819	.526	.586	.879	.930	.823	.546	.593	.868

Note. F1 = Factor 1; F2 = Factor 2; F3 = Factor 3; F4 = Factor 4.

5 Discussion

Some scholars argue that the concept of LLS is outdated and that the field should focus more on self-regulation (Dörnyei, 2005; Habók et al., 2022). However, Pawlak (2021) finds that research on LLS remains vibrant and continues to shed light on how learners acquire additional languages. There has been a noticeable improvement in students' language learning when LLS are applied. Using different LLS and tailoring them to individual preferences can enhance learning outcomes (Yoong & Hashim, 2023). In a rapidly globalising society, students must stay flexible and adopt new strategies, which may include technological tools (James et al., 2019). The skills gained through these innovative methods have helped students manage challenges effectively.

This study aimed to adapt the SILL questionnaire for secondary school students in the Czech Republic. The final model identified two factors: direct (11 items) and indirect strategies (6 items). This indicates that the original six-factor structure by Oxford (1989) was not maintained. Hsiao and Oxford (2003), in their follow-up research using CFA, suggested that either a two-factor or a six-factor model would be suitable, aligning with our findings of a two-factor solution. While the SILL has been adapted in various studies over time, results have varied significantly (Park, 2011). Some research validated the original six-factor structure (Demirel, 2009; Wu & Chang, 2018), whereas others did not (Park, 2011; Yesilbursa & Ipek, 2013). Both EFA and CFA were used in these studies, but the inconsistency prompted us to employ both methods to assess construct validity.

In the two-factor model, items associated with the first factor (direct strategies) mainly reflect the cognitive aspects of LLS. Cognitive strategies involve how learners process and transform language. Vlčková (2005) found that cognitive strategies were the most effective among those identified by R. Oxford. This study also confirmed that cognitive strategies predominate, supporting Dörnyei's (2005) critique that Oxford's classification incorrectly separates memory strategies from cognitive ones. Cognitive functions encompass attention, judgment, evaluation, problem-solving, understanding, and language production, with memory being a key component (Wood & Rutterford, 2006). Prior research with high school and college students showed that cognitive strategies are used more frequently than metacognitive strategies, indicating their dominant role (Chamot et al., 1987; O'Malley et al., 1985).

Within EFA, the .40–.30–.20 rule established by Howard (2016) was strictly followed. This rule states that factor loadings demonstrate whether items are suitable indicators of their respective scales. All 17 items in the two-factor model had factor loadings above .40, but this was not true for the three-factor model. Generally, higher (stronger) factor loadings make it easier to interpret the factor's direction (Field, 2024). The first items (I1 to I4) of the first factor in the two-factor model displayed the highest loadings (.746 to .802), emphasising communication skills. This aligns with language teaching, especially English. Although students knew grammatical rules and had extensive vocabularies, they still

lacked communication skills (Saks & Leijen, 2016). Research on English as a lingua franca (ELF) addresses this gap by focusing on communicative functions rather than linguistic forms to promote mutual understanding (Chan, 2021).

Within the CFA, several items were linked, chiefly the English reading and writing domains of I5 and I6 (such as writing notes, messages, letters, or reports in English, and reading English for pleasure). Esmaeili's (2002) study demonstrated that thematic connections between reading and writing improved both the processes and outcomes of students' writing.

The second connection was between I5 and I9 (I write notes, messages, letters, or reports in English; I encourage myself to speak English even when I am afraid of making a mistake). Ergün Öcel (2023) suggests that educators should focus on developing skills that enhance oral communication. A good starting point is to support this through English writing. Namanziandost et al. (2018) highlight how writing can significantly help deepen and improve spoken language. Their study showed that, in an oral post-test, students made notable gains in grammar usage, vocabulary breadth, and syntactic structure production.

The third link focused on items I6 and I9 (I read English for pleasure; I encourage myself to speak English even when I am afraid of making a mistake). Jia's (2023) research found a strong correlation between reading and speaking in English. Reading in English affects pronunciation, vocabulary, and comprehension of the language.

The second-to-last link for the first factor included items I6 and I7 (I read English for pleasure; I look for opportunities to read as much as possible in English). Reading is an essential part of learning English. Through reading, students gain new vocabulary, develop a sense of coherent speech, and understand logical structuring in writing. Ansarey's (2016) study shows that 80% of respondents felt that reading English texts enhanced their speaking skills.

The final fifth link was for items I6 and I8 (I read for pleasure in English; I watch English-language TV shows or go to English-language movies). Ennemoser and Schneider (2007) found in their longitudinal study that watching educational programs was positively correlated with reading performance, whereas entertainment programs were negatively correlated. Individuals who spent an average of 117 minutes per day watching television showed lower reading progress than individuals who spent an average of 35-69 minutes per day watching programs.

The second factor (indirect strategies) comprises supportive activities in the learning process (Oxford, 1990). Indirect strategies encompass metacognitive strategies, which involve understanding one's own cognitive processes (knowledge of cognition) and the ability to control them (regulation of cognition). Thus, cognitive and metacognitive strategies cannot be strictly separated (Hrbáčková, 2011). In the second factor, only one link emerged between items I14 and I17 (I review English lessons regularly; I say or write new English words several times). These items may be closely related because of their predominantly cognitive character. Nevertheless, the analysis in this study placed them in the second factor. This placement can be interpreted as reflecting metacognitive knowledge. Metacognitive knowledge comprises contextual knowledge, which enables learners to select appropriate strategies and adapt their learning to the changing demands of tasks. This outcome may be interpreted as reflecting metacognitive knowledge. For example, a student knows how and which information to revise to achieve success in an exam (Švec, 2005).

To further evaluate the validity of the two-factor model, a correlation analysis was performed. A moderately strong relationship emerged between the factors, supporting the earlier observation of the connection between cognitive and metacognitive strategies. Flavell (1979) supports this perspective, suggesting that cognitive and metacognitive strategies are complementary rather than entirely distinct. Metacognition helps students manage their cognitive processes more effectively, such as analysing problems, planning, and monitoring the success of their strategies. Consequently, this enhances learning and problem-solving abilities. From a broader standpoint, these results can also be viewed through the lens of social pedagogy. Learning strategies can be seen as indicators of how

students handle school demands and regulate their learning, strongly linked to both academic achievement and the risk of failure. The validated SILL-CZ instrument thus functions not only as a research tool but also as a practical diagnostic resource for teachers and school social pedagogues, aiding early detection of learning difficulties and guiding targeted interventions to prevent academic failure.

5.1 Limitations and future directions

The SILL-CZ questionnaire is designed for secondary school students, so its applicability is limited to that age group. The findings might not generalise to other age groups. Additionally, because the data rely on students' self-assessments, they could be biased. The questionnaire might also fail to fully capture the evolving and context-dependent nature of learning strategies in real-world settings.

Future research should place greater emphasis on integrating questionnaires with qualitative methods such as learner diaries, classroom observations, and introspective techniques to more effectively capture the complexity of strategy use. Pawlak (2021) emphasises that strategies are context-dependent and constantly evolving, influenced by factors such as task type, motivation, and emotional states. Consequently, students may employ different strategies when preparing for regular lessons versus secondary school leaving exams.

5 Conclusion

The aim of this study was to describe the adaptation process of the SILL, Version for Speakers of Other Languages Learning English questionnaire for the Czech population of secondary school students. Factor analysis revealed that the original six-factor model (Oxford, 1989) was not confirmed in the Czech context. Instead, a two-factor structure, consisting of direct and indirect strategies, proved to be a better fit. The results also highlighted the predominant role of direct strategies, which are primarily associated with the cognitive domain.

The study, from a theoretical standpoint, advances the discussion on the validity of the SILL and the nature of language learning strategies, which are recognised as situation-dependent and dynamic (see Pawlak, 2021). Additionally, it contributes to the broader debate on the appropriate classification of strategies in the Czech context.

In practice, the findings can help teachers and curriculum designers enhance their work on language learning strategies, especially given recent changes to the national secondary school leaving exam. A key practical contribution is a simple two-factor inventory that offers educators a quick diagnostic tool and provides students with a clear guide for self-monitoring their use of strategy. This improves its usefulness in classroom settings, particularly in secondary education and exam preparation, where supporting students involves both language acquisition and the development of self-regulation skills. Additionally, from a social pedagogical perspective, these findings can aid in early identification of students at risk of school failure and enable targeted interventions to promote better learning regulation.

Utilising a two-factor model that includes both direct and indirect learning strategies enhances English language acquisition, self-regulation, and academic achievement among Czech secondary school students.

Generative AI disclosure

No Generative AI was involved in preparing this manuscript.

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Supplementary Table 1

Factor loadings and communalities for the three-factor solution in the EFA of the SILL.

Item	Factor			h^2	
	1	2	3		
I1	I try to find as many ways as I can to use my English.	.822		.677	
I2	I start conversations in English.	.785		.630	
I3	I write notes, messages, letters, or reports in English.	.754		.592	
I4	I read English for pleasure.	.745		.594	
I5	I look for opportunities to read as much as possible in English.	.729		.568	
I6	I watch English-language TV shows or go to movies in English.	.667		.509	
I7	I read English without looking up every new word.	.640		.505	
I8	I try to learn about the culture of English speakers.	.583		.348	
I9	I plan my schedule so I will have enough time to study English.		.653	.468	
I10	I write down my feelings in a language learning diary.		.509	.277	
I11	I use rhymes to remember new English words.		.399	.165	
I12	I talk to someone else about how I feel while learning English.		.389	.193	
I13	I use flashcards to remember new English words.		.370	.157	
I14	I physically act out new English words.		.355	.130	
I15	I find the meaning of an English word by dividing it into parts that I understand.			.559 .336	
I16	I notice when I am tense or nervous while studying or using English.			.525 .288	
I17	If I do not understand something in English, I ask the other person to slow down or say it again.			.449 .316	
I18	I remember new English words or phrases by remembering their location on the page, on the board, or on a street sign.			.433 .207	
				Total	
	Cronbach's α	.898	.601	.546	.802
	McDonald's ω	.900	.591	.543	.802

Note. Extraction method: Principal Axis Factoring. Rotation Method: Oblimin with Kaiser Normalisation. I = item; h^2 = communalities. F1 = Factor 1; F2 = Factor 2; F3 = Factor 3; F4 = Factor 4.

Supplementary Table 2

Factor loadings and communalities for the four-factor solution in the EFA of the SILL.

Item	Factor				h^2	
	1	2	3	4		
I1	I try to find as many ways as I can to use my English.	.818				.678
I2	I start conversations in English.	.789				.641
I3	I write notes, messages, letters, or reports in English.	.764				.615
I4	I try to talk like a native English speaker.	.757				.595
I5	I read English for pleasure.	.755				.597
I6	I ask questions in English.	.746				.568
I7	I look for opportunities to read as much as possible in English.	.723				.567
I8	I watch English-language TV shows or go to movies in English.	.672				.508
I9	I look for people I can talk to in English.	.665				.582
I10	I read English without looking up every new word.	.655				.536
I11	I pay attention when someone is speaking English.	.655				.533
I12	I use the English words I know in different ways.	.630				.506
I13	I try to learn about the culture of English speakers.	.576				.348
I14	I try to find out how to be a better learner of English.		.734			.566
I15	I have clear goals for improving my English skills.		.734			.592
I16	I plan my schedule so I will have enough time to study English.		.652			.490
I17	I think about my progress in learning English.		.647			.468
I18	I review English lessons regularly.		.572			.361
I19	I notice when I am tense or nervous while studying or using English.			.502		.286
I20	I make up new words if I do not know the right ones in English.			.473		.240
I21	When I can't think of a word during an English conversation, I use gestures.			.425		.187
I22	I write down my feelings in a language learning diary				.574	.341
I23	I use rhymes to remember new English words.				.523	.278
I24	I physically act out new English words.				.475	.239
						Total
	Cronbach's α	.929	.819	.526	.586	.903
	McDonald's ω	.930	.823	.546	.583	.905

Note. Extraction method: Principal Axis Factoring. Rotation Method: Oblimin with Kaiser Normalisation. I = item; h^2 = communalities. F1 = Factor 1; F2 = Factor 2; F3 = Factor 3; F4 = Factor 4.

Supplementary Table 3

Descriptive statistics for the 17 final items of the two-factor SILL-CZ solution.

Factor	Item	M	SD	Skew.	Kurt.	
Direct Strategies	I1	I try to find as many ways as I can to use my English.	3.21	1.323	.140	-1.079
	I2	I try to talk like a native English speaker.	3.27	1.408	.240	-1.199
	I3	I ask questions in English.	3.16	1.406	.138	-1.230
	I4	I start conversations in English.	2.97	1.471	.048	-1.357
	I5	I write notes, messages, letters, or reports in English.	2.84	1.515	.218	-1.398
	I6	I read English for pleasure.	3.14	1.464	.119	-1.345
	I7	I look for opportunities to read as much as possible in English.	2.63	1.387	.355	-1.107
	I8	I watch English-language TV shows or go to English-language movies.	3.81	1.376	.783	-.757
	I9	I encourage myself to speak English even when I am afraid of making a mistake.	3.40	1.343	.435	-.950
	I10	I try to learn about the culture of English speakers.	3.03	1.453	.015	-1.337
	I11	I try not to translate word-for-word.	3.39	1.297	.312	-.992
Indirect Strategies	I12	I plan my schedule so I will have enough time to study English.	1.95	1.110	1.065	-.375
	I13	I try to find out how to be a better learner of English.	2.96	1.335	.014	-1.106
	I14	I review English lessons often.	2.16	1.093	.590	-.537
	I15	I look for words in my own language that are similar to new English words.	2.39	1.307	.549	-.829
	I16	I give myself a reward or treat when I do well in English.	2.47	1.394	.495	-1.044
	I17	I say or write new English words several times.	2.64	1.392	.303	-1.187

Note. I = Item; M = mean; SD = standard deviation; Skew. = skewness; Kurt. = kurtosis.

Supplementary Figure 1
 Two-factor CFA model of the SILL-CZ.

